

Time Series: Dynamic Models and Autocorrelation

Modeling Temporal Dependence in Economic Data

Jake Anderson

March 3, 2026

Outline

- 1 What Makes Time Series Different
- 2 Stationarity
- 3 Autocorrelation Function (ACF)
- 4 AR(p) Models
- 5 ARDL Models
- 6 Serial Correlation
- 7 Forecasting
- 8 Summary

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\implies The i.i.d. assumption fails. We need new tools that account for **temporal dependence**.

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- 2 The **order** of observations contains information we should exploit
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This lecture covers:

- How to characterize temporal dependence (stationarity, ACF)
- Models that exploit it (AR, ARDL)
- How to detect it in regression residuals (Breusch-Godfrey)
- How to forecast with it

Covariance Stationarity: Definition

A time series $\{Y_t\}$ is **covariance stationary** if:

- 1 **Constant mean:** $E(Y_t) = \mu$ for all t
- 2 **Constant variance:** $\text{Var}(Y_t) = \sigma^2$ for all t
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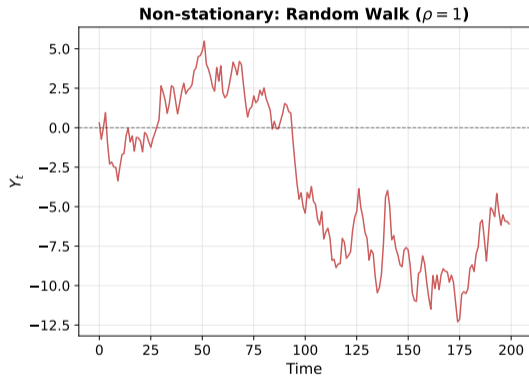
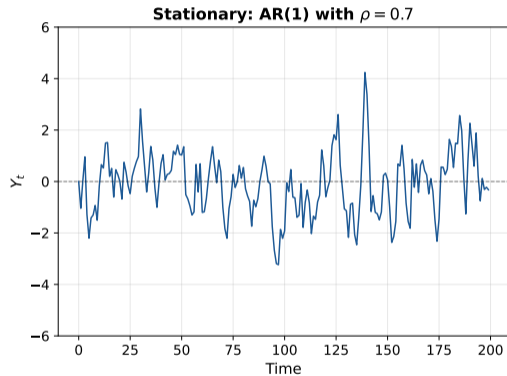
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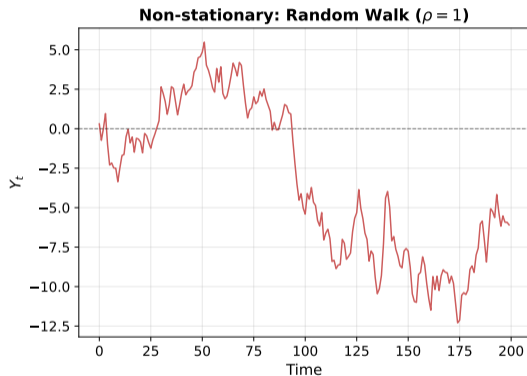
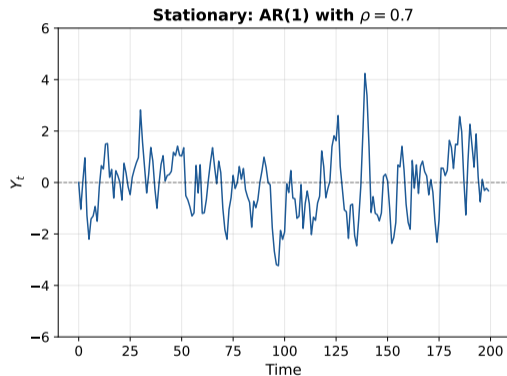
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⇒ Stationarity ensures that estimated statistical properties are meaningful and stable.

Stationary vs Non-stationary: Visual Comparison



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Left: mean-reverting; always pulled back toward zero. **Right:** wanders without bound; no tendency to return.

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Example: Stock prices are approximately random walks. That is why financial economists work with *returns* (which are approximately stationary) rather than price levels.

Measuring Temporal Dependence: The ACF

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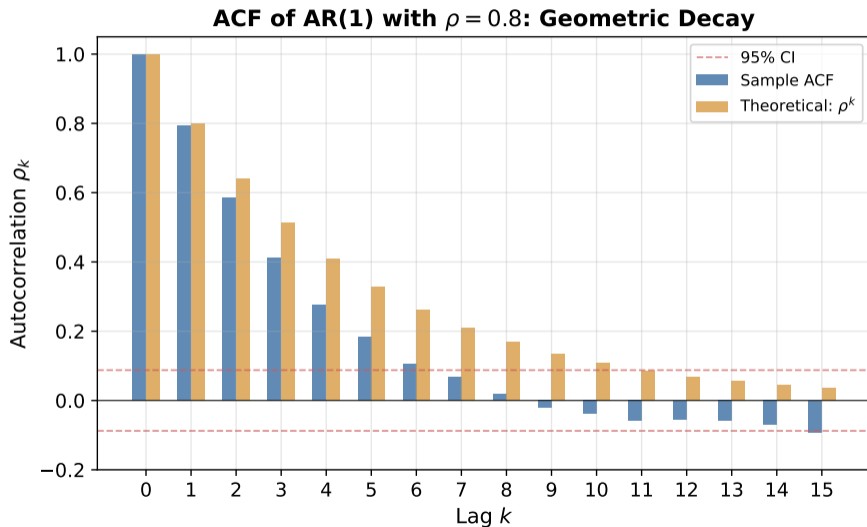
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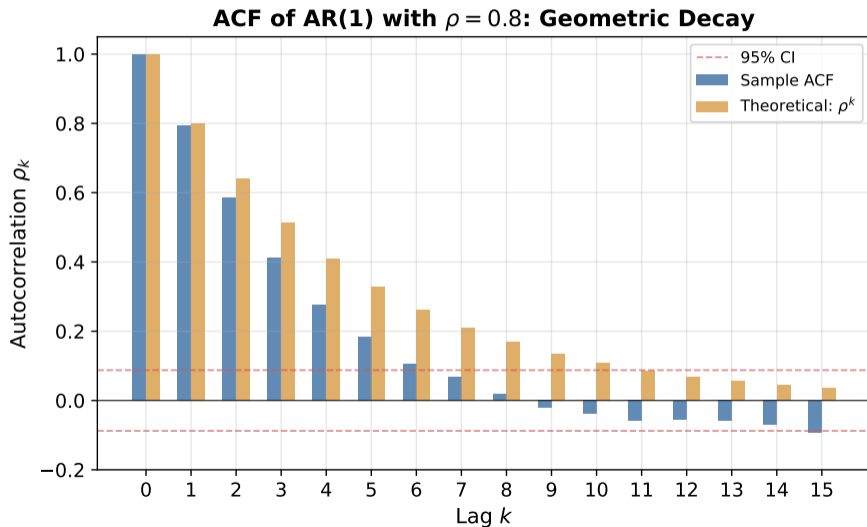
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The **sample autocorrelation**:

$$r_k = \frac{\sum_{t=k+1}^T (Y_t - \bar{Y})(Y_{t-k} - \bar{Y})}{\sum_{t=1}^T (Y_t - \bar{Y})^2}$$





Autoregressive Models: AR(p)

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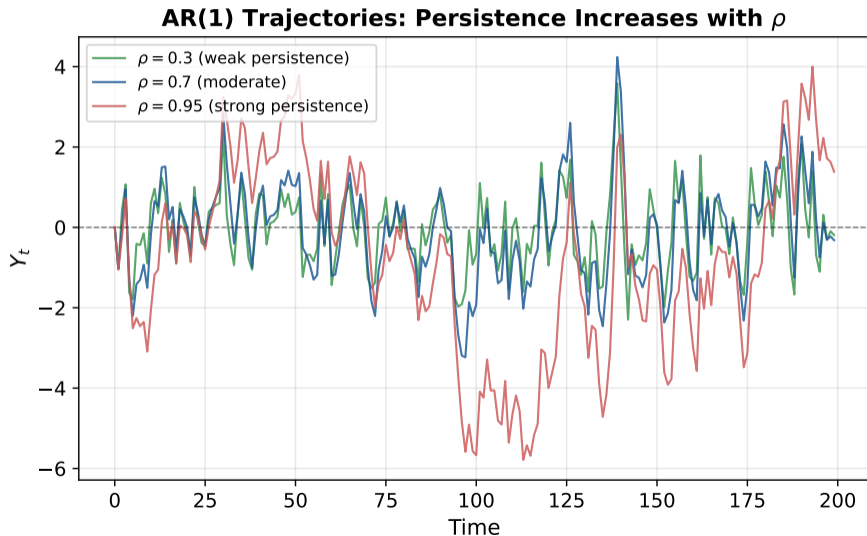
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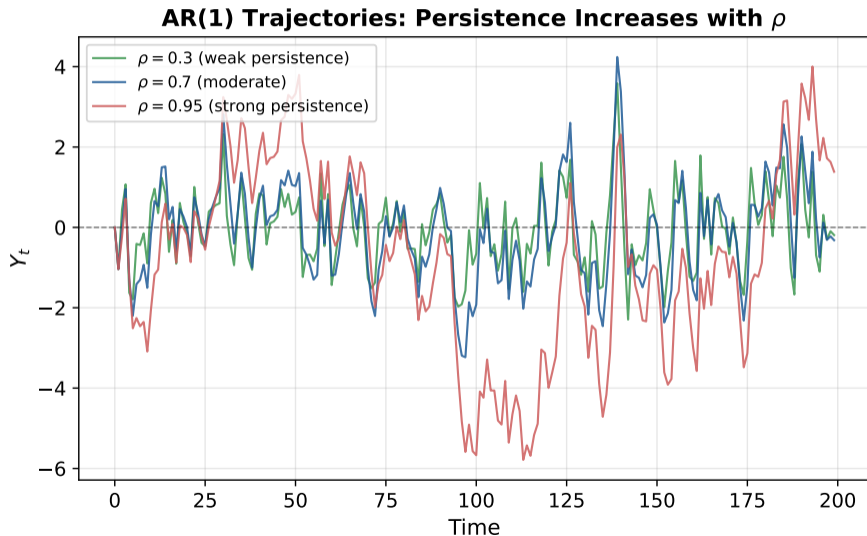
AR(2): $Y_t = \delta + \theta_1 Y_{t-1} + \theta_2 Y_{t-2} + v_t$

- Allows richer dynamics: oscillations, humps

How Persistence Changes with ρ



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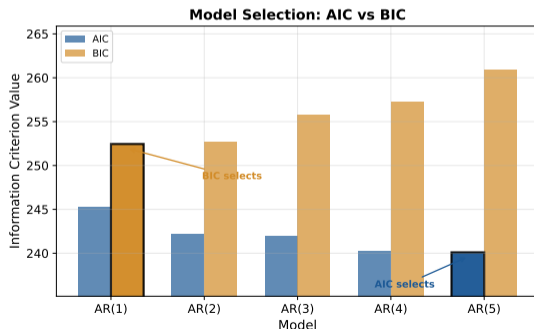
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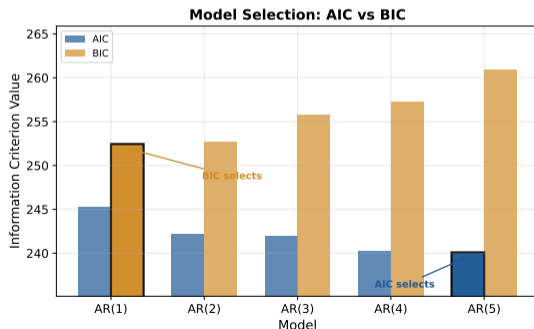


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Autoregressive Distributed Lag: ARDL(p, q)

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⇒ Current inflation depends on: (1) its own recent history, and (2) current *and* past unemployment changes.

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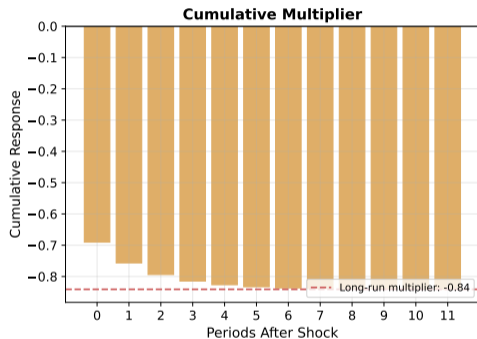
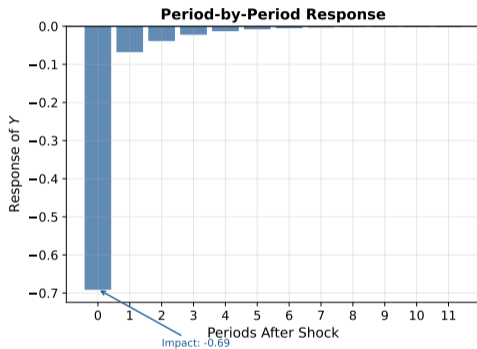
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\implies The long-run effect is larger (in absolute value) than the impact effect whenever $|\theta_1| > 0$.

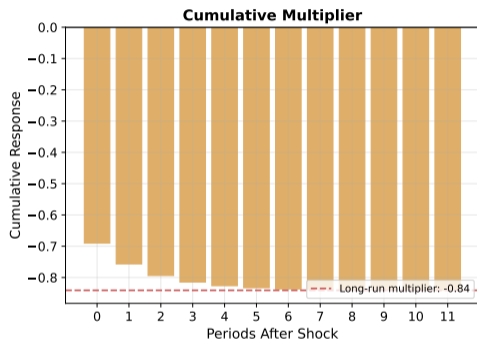
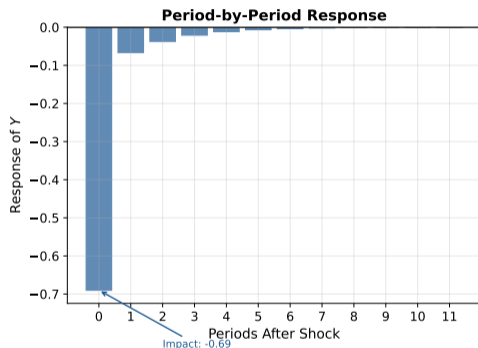
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- Impact multiplier: -0.69
- Interim: $-0.69 + 0.32 = -0.37$
- Long-run: $\frac{-0.69+0.32}{1-0.56} = -0.84$

Multiplier Practice Problem

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\implies The long-run effect (-0.84) exceeds the impact (-0.69) because the AR term propagates the shock.

Serial Correlation in Regression Residuals

Suppose we estimate:

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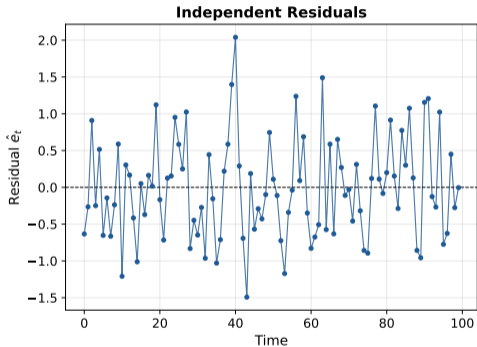
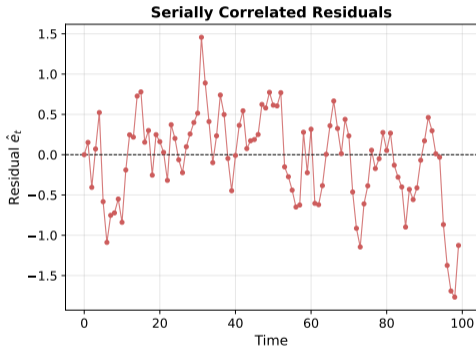
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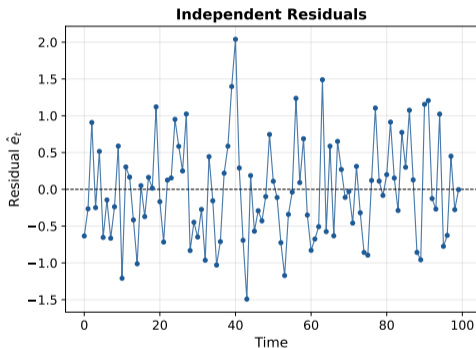
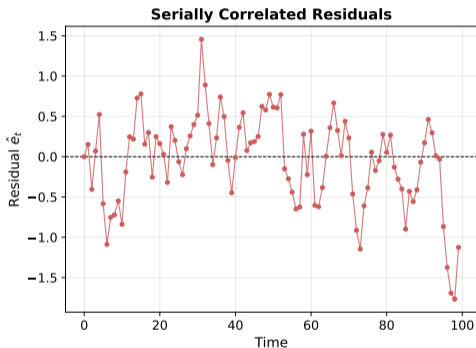
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\implies The coefficients are fine, but our **inference** (tests, CIs) is unreliable.

What Serially Correlated Residuals Look Like



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Left: runs of positive then negative residuals \implies positive autocorrelation. **Right:** residuals bounce randomly around zero \implies no serial correlation.

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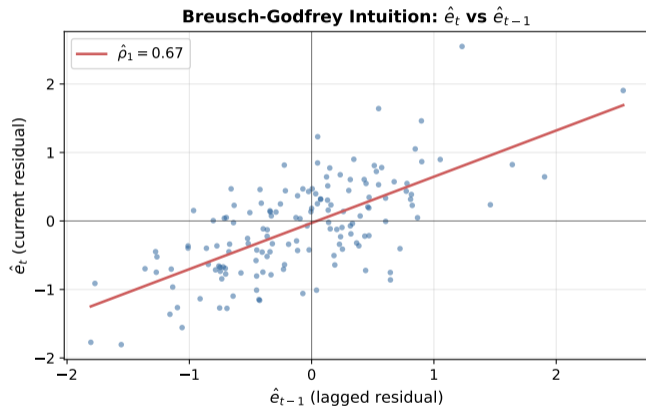
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Important: the auxiliary regression *includes all original regressors*. This is what makes BG valid even with lagged dependent variables (unlike Durbin-Watson).

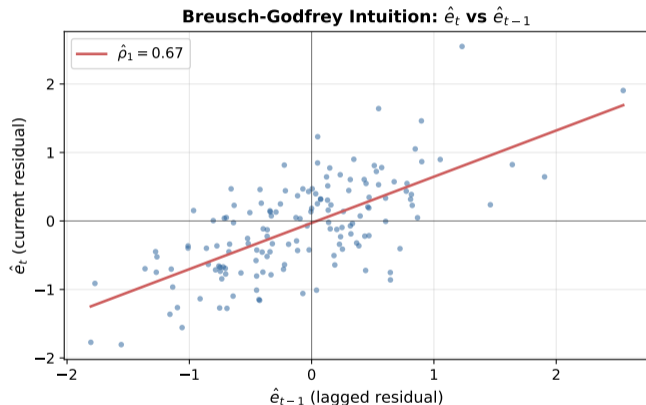
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The BG auxiliary regression checks whether $\hat{\rho}_1$ (the slope in this scatter) is significantly different from zero, *after controlling for the original regressors*.

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⇒ Each forecast builds on previous forecasts, compounding uncertainty.

Forecast Practice Problem

An AR(2) model for quarterly inflation gives:

$$\hat{\delta} = 0.4523, \quad \hat{\theta}_1 = 0.6234, \quad \hat{\theta}_2 = 0.2145$$

with $INF_T = 2.5$ and $INF_{T-1} = 3.0$.

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Solution:

$$\begin{aligned}\widehat{\text{INF}}_{T+1} &= 0.4523 + 0.6234(2.5) + 0.2145(3.0) \\ &= 0.4523 + 1.5585 + 0.6435 \\ &= 2.65\end{aligned}$$

Forecast Intervals

As we forecast further ahead, uncertainty grows:

$$\hat{Y}_{T+h} \pm t_c \cdot \hat{\sigma}_h$$

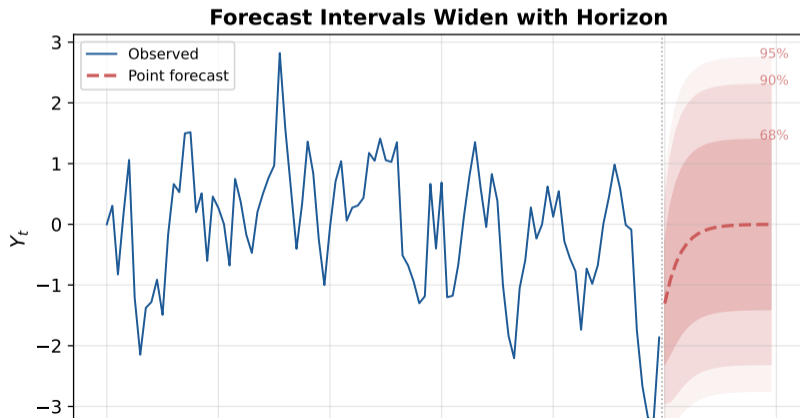
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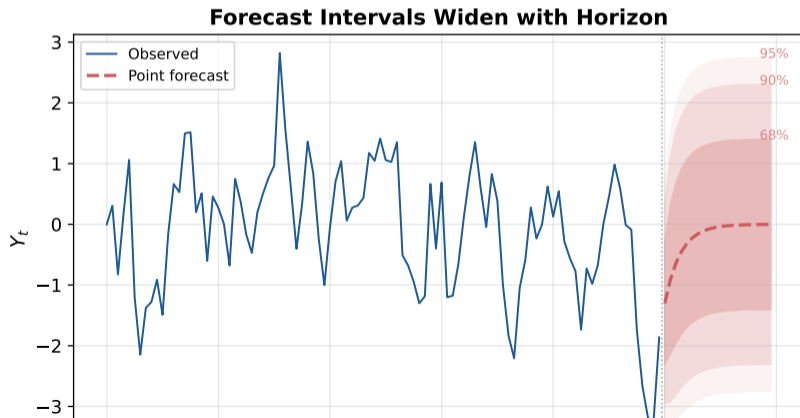


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⇒ Always evaluate forecasts out-of-sample when the goal is prediction.

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- 7 **Forecasting**: iterate AR/ARDL forward; confidence bands widen with horizon. Evaluate out-of-sample.

Thank you!
jakeanderson@g.ucla.edu