

# Time Series: Dynamic Models and Autocorrelation

## Modeling Temporal Dependence in Economic Data

Jake Anderson

May 16, 2026

# Outline

- 1 What Makes Time Series Different
- 2 Stationarity
- 3 Autocorrelation Function (ACF)
- 4 AR( $p$ ) Models
- 5 ARDL Models
- 6 Serial Correlation
- 7 Forecasting
- 8 Summary

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$\implies$  The i.i.d. assumption fails. We need new tools that account for **temporal dependence**.

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When observations are correlated over time:

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- 2 The **order** of observations contains information we should exploit
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This lecture covers:

- How to characterize temporal dependence (stationarity, ACF)
- Models that exploit it (AR, ARDL)
- How to detect it in regression residuals (Breusch-Godfrey)
- How to forecast with it

# Covariance Stationarity: Definition

A time series  $\{Y_t\}$  is **covariance stationary** if:

- 1 **Constant mean:**  $E(Y_t) = \mu$  for all  $t$
- 2 **Constant variance:**  $\text{Var}(Y_t) = \sigma^2$  for all  $t$
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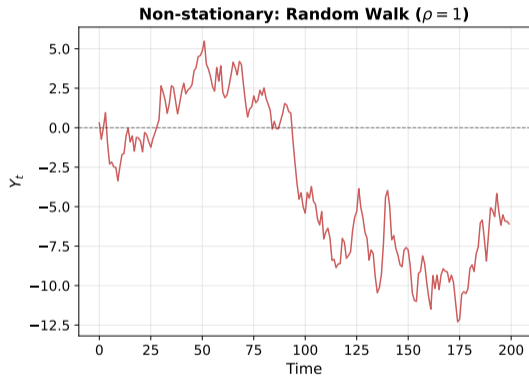
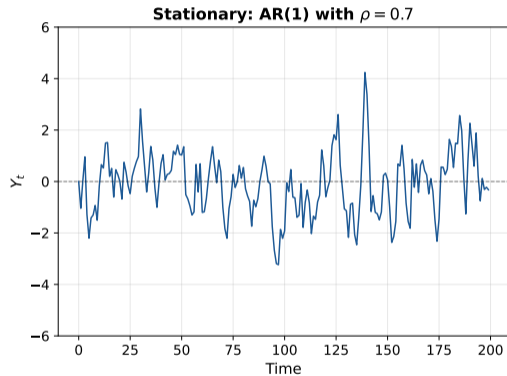
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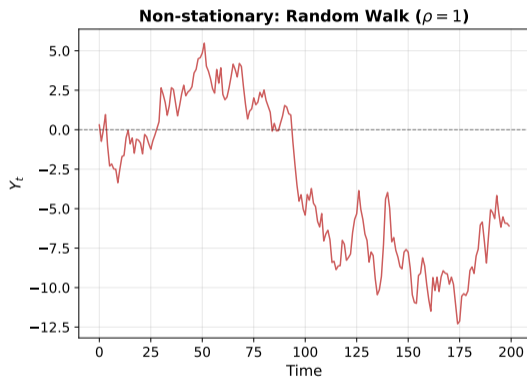
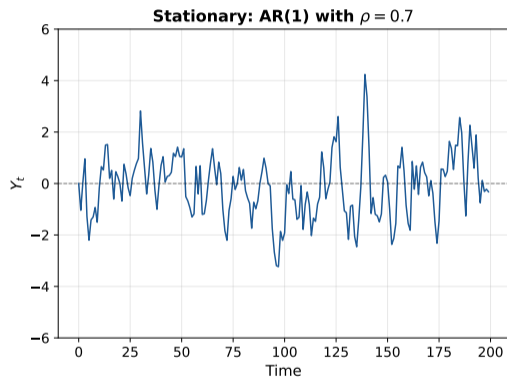
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⇒ Stationarity ensures that estimated statistical properties are meaningful and stable.

# Stationary vs Non-stationary: Visual Comparison



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**Left:** mean-reverting; always pulled back toward zero. **Right:** wanders without bound; no tendency to return.

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**Example:** Stock prices are approximately random walks. That is why financial economists work with *returns* (which are approximately stationary) rather than price levels.

# Measuring Temporal Dependence: The ACF

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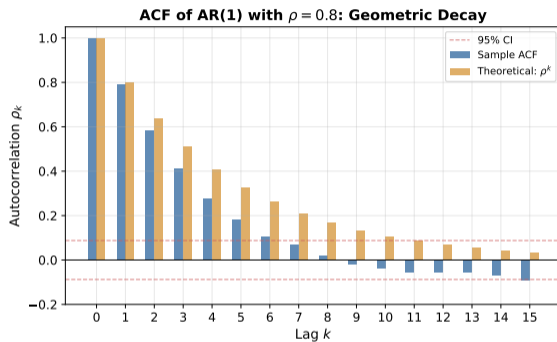
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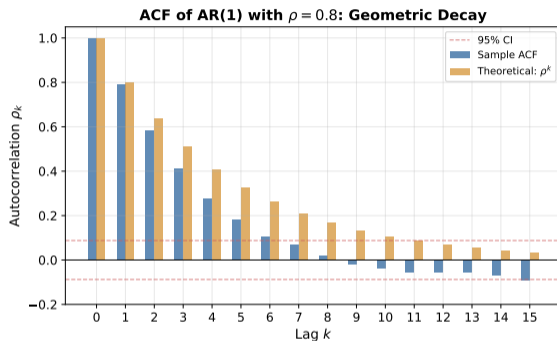
The **sample autocorrelation**:

$$r_k = \frac{\sum_{t=k+1}^T (Y_t - \bar{Y})(Y_{t-k} - \bar{Y})}{\sum_{t=1}^T (Y_t - \bar{Y})^2}$$

# Reading an ACF Plot



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## How to read it:

- Bars beyond the dashed lines = statistically significant autocorrelation
- For an AR(1), the ACF decays **geometrically**:  $\rho_k = \rho^k$
- If residuals from a model show significant ACF spikes  $\implies$  the model is missing dynamics

## Autoregressive Models: AR( $p$ )

An **AR( $p$ ) model** says today's value depends on its own past  $p$  values:

$$Y_t = \delta + \theta_1 Y_{t-1} + \theta_2 Y_{t-2} + \cdots + \theta_p Y_{t-p} + v_t$$

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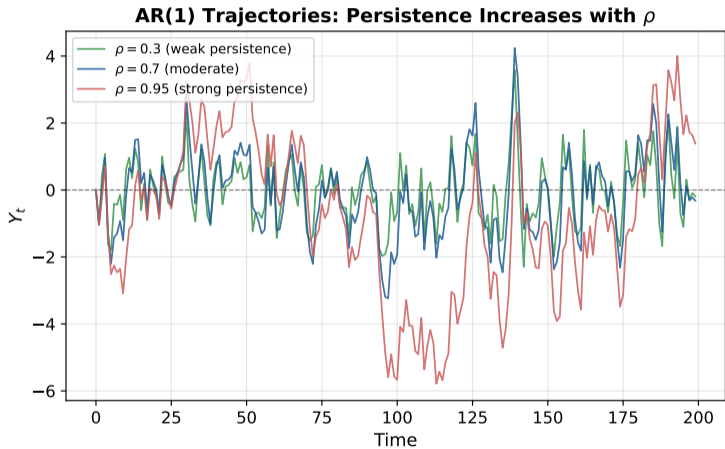
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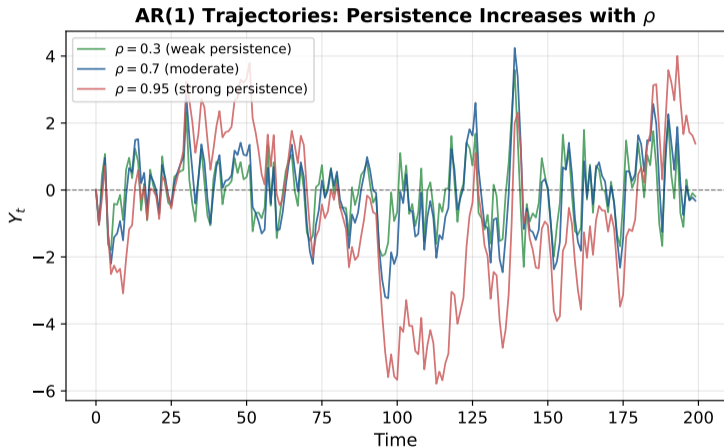
**AR(2):**  $Y_t = \delta + \theta_1 Y_{t-1} + \theta_2 Y_{t-2} + v_t$

- Allows richer dynamics: oscillations, humps

# How Persistence Changes with $\rho$



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Same shocks, different  $\rho$ . Higher  $\rho \implies$  slower mean reversion, longer “memory.”

## Model Selection: AIC and BIC

How many lags? Use **information criteria**:

$$\text{AIC} = \ln(\hat{\sigma}^2) + \frac{2K}{T}, \quad \text{BIC} = \ln(\hat{\sigma}^2) + \frac{K \ln(T)}{T}$$

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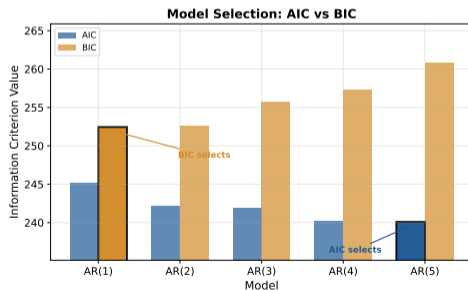
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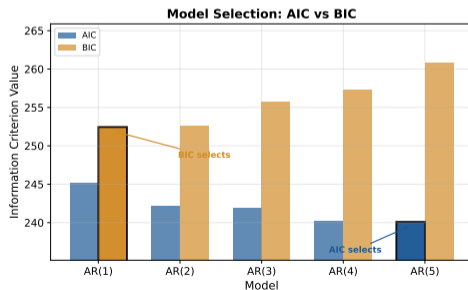


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When AIC and BIC disagree, BIC is typically preferred for consistent selection.

# Autoregressive Distributed Lag: ARDL( $p, q$ )

An **ARDL** model includes lags of both  $Y$  and an explanatory variable  $X$ :

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where  $\text{INF}$  = inflation,  $\text{DU}$  = change in unemployment.

$\implies$  Current inflation depends on: (1) its own recent history, and (2) current *and* past unemployment changes.

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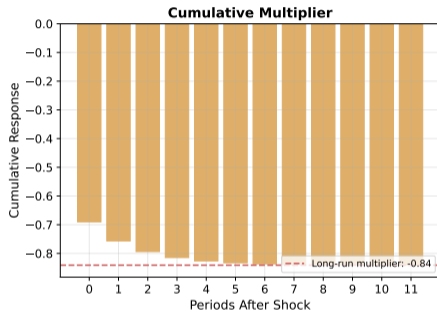
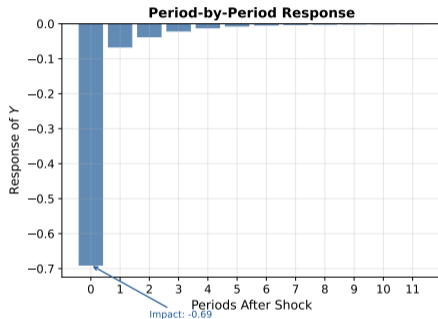
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$\implies$  The long-run effect is larger (in absolute value) than the impact effect whenever  $|\theta_1| > 0$ .

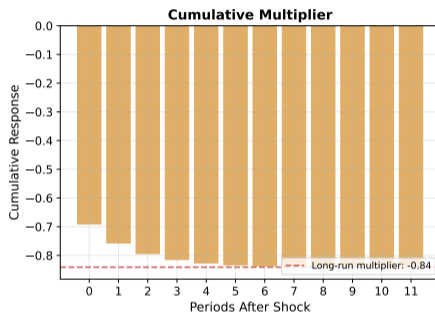
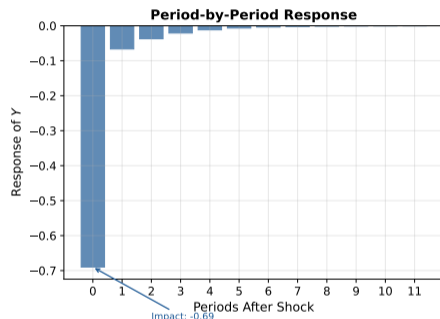
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- Impact multiplier:  $-0.69$
- Interim:  $-0.69 + 0.32 = -0.37$
- Long-run:  $\frac{-0.69+0.32}{1-0.56} = -0.84$

# Multiplier Practice Problem

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$\implies$  The long-run effect ( $-0.84$ ) exceeds the impact ( $-0.69$ ) because the AR term propagates the shock.

# Serial Correlation in Regression Residuals

Suppose we estimate:

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## Consequences:

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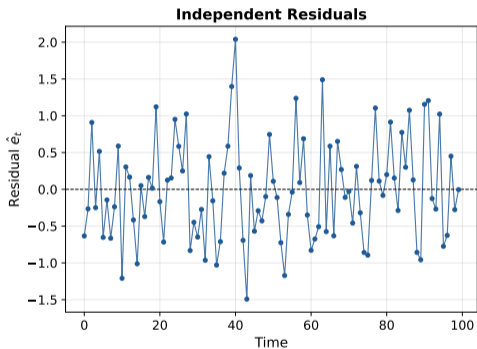
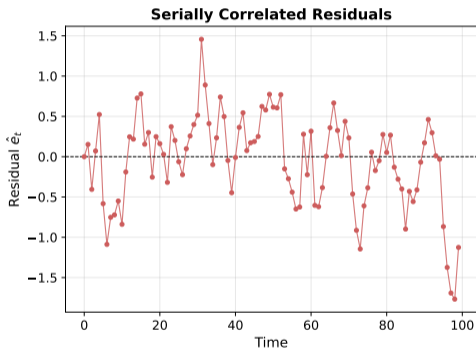
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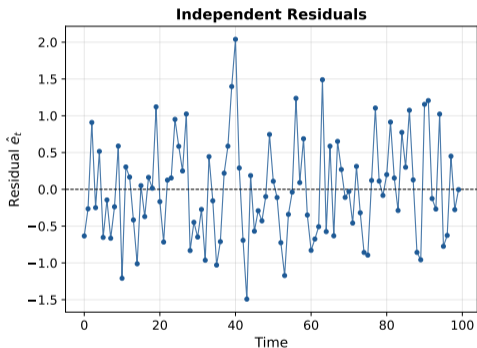
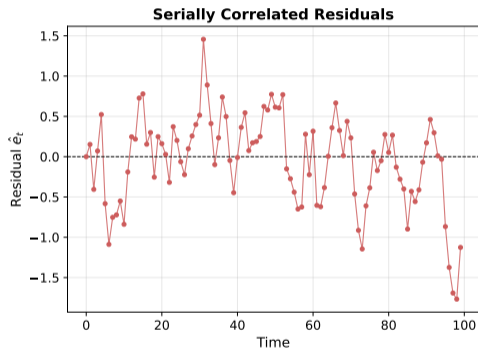
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$\implies$  The coefficients are fine, but our **inference** (tests, CIs) is unreliable.

# What Serially Correlated Residuals Look Like



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**Left:** runs of positive then negative residuals  $\implies$  positive autocorrelation. **Right:** residuals bounce randomly around zero  $\implies$  no serial correlation.

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Tests whether residuals are autocorrelated up to order  $q$ .

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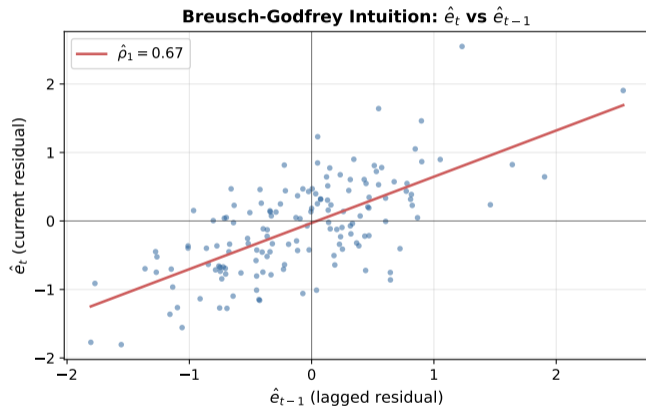
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**Important:** the auxiliary regression *includes all original regressors*. This is what makes BG valid even with lagged dependent variables (unlike Durbin-Watson).

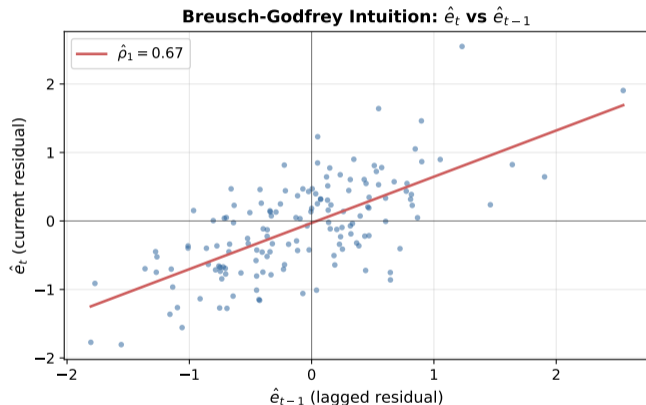
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The BG auxiliary regression checks whether  $\hat{\rho}_1$  (the slope in this scatter) is significantly different from zero, *after controlling for the original regressors*.

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Plug in the most recent observed values and estimated coefficients.

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⇒ Each forecast builds on previous forecasts, compounding uncertainty.

## Forecast Practice Problem

An AR(2) model for quarterly inflation gives:

$$\hat{\delta} = 0.4523, \quad \hat{\theta}_1 = 0.6234, \quad \hat{\theta}_2 = 0.2145$$

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**Compute the one-step-ahead forecast:**

**Solution:**

$$\begin{aligned}\widehat{\text{INF}}_{T+1} &= 0.4523 + 0.6234(2.5) + 0.2145(3.0) \\ &= 0.4523 + 1.5585 + 0.6435 \\ &= 2.65\end{aligned}$$

# Forecast Intervals

As we forecast further ahead, uncertainty grows:

$$\hat{Y}_{T+h} \pm t_c \cdot \hat{\sigma}_h$$

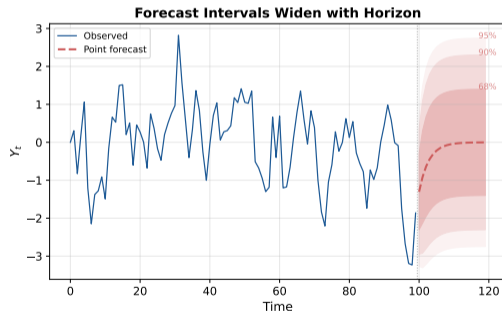
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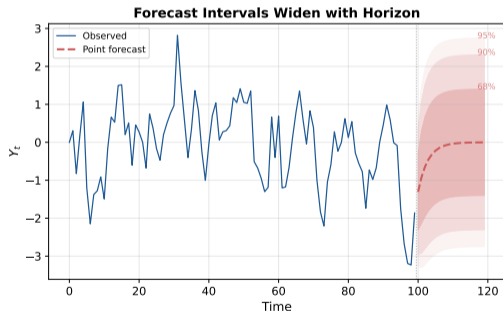


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⇒ Short-horizon forecasts are relatively precise. Long-horizon forecasts revert toward the unconditional mean with wide bands.

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⇒ Always evaluate forecasts out-of-sample when the goal is prediction.

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- 7 **Forecasting**: iterate AR/ARDL forward; confidence bands widen with horizon. Evaluate out-of-sample.

Thank you!  
jakeanderson@g.ucla.edu